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1	Estimating trophic levels and trophic magnification factors
2	using Bayesian inference
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### 24 ABSTRACT

Food web biomagnification is increasingly assessed by estimating trophic magnification 25 factors (TMF) where solvent (often lipid) normalised contaminant concentration is regressed 26 onto trophic level, and TMFs are represented by the slope of the relationship. In TMF 27 regressions, the uncertainty in the contaminant concentrations is appreciated, whereas the 28 trophic levels are assumed independent and not associated with variability or uncertainty 29 30 pertaining to e.g. quantification. In reality, the trophic levels may vary due to measurement error in stable isotopes of nitrogen ( $\delta^{15}N$ ) of each sample, in  $\delta^{15}N$  in selected reference 31 32 baseline trophic level, and in the enrichment factor of  $\delta^{15}N$  between two trophic levels ( $\Delta N$ ), which are all needed to calculate trophic levels. The present study used a Markov Chain 33 Monte Carlo method, with knowledge about the food web structure, which resulted in a 34 dramatic increase in the precision in the TMF estimates. This also lead to a better 35 36 understanding of the uncertainties in bioaccumulation measures; instead of using point estimates of TMF, the uncertainty can be quantified (i.e. TMF >1, namely positive 37 biomagnification, with an estimated X % probability). 38 39 Keywords: biomagnification, trophic level, food web, contaminants 40

#### 42 INTRODUCTION

Recent reviews and studies have suggested the implementation of trophic relations in the 43 assessment guidelines of contaminant accumulation<sup>1-4</sup>. This includes evaluating the 44 bioaccumulation potential of contaminants by quantifying their magnification through diet, 45 either by specific predator-prey relations (biomagnification factor - BMF) or as an average 46 factorial change from one trophic level to the next in a specified food web (trophic 47 48 magnification factor –TMF; previously also referred to as Food Web Magnification Factor). Whereas the BMF is the ratio of contaminant concentration between predator and prey 49 (BMF=C<sub>PREDATOR</sub>/C<sub>PREY</sub>), the TMF is estimated by regressing the contaminant concentrations 50 51 in representatives of a food web onto their relative trophic positions, and the TMF is the slope of the regression line <sup>3,5,6</sup>. Although the TMF is currently recognized as the most realistic 52 quantitative measure of food web accumulation of contaminants<sup>1,4</sup>, several issues remain 53 54 regarding scientific understanding, feasibility of test protocols, and thus regulatory acceptance<sup>7,8</sup>. One of the greater challenges is to obtain a better understanding of the 55 56 variability in TMF estimates and whether this variability comes about through natural 57 variation in relevant processes or uncertainties surrounding our knowledge of them, or if it is the result of measurement errors, poorly defined concepts and statistical analyses. Despite 58 59 this, the European Community Regulation on chemicals and their safe use (REACH) recently amended to Annex XIII that accumulation of chemicals from the diet (BMF) and in the food 60 web (TMF) could be used in the weight of evidence assessment of the chemical as a 61 contaminant of concern due to bioaccumulation (REACH, Annex XIII<sup>9</sup>). 62

63

64 The trophic level of a species reflects its approximate feeding position in a food web,
65 where primary producers (plants/algae) constitute the first trophic level, followed by primary
66 consumers (herbivore) on the second trophic level, secondary and tertiary consumers

(carnivore) on the third and fourth trophic level, and so on. However, the simple concept of 67 unidirectional linear food chains rarely apply to natural ecosystems, where more complex 68 network models more appropriate describe the food webs<sup>10</sup>. Thus, the feeding position of a 69 species is not an integer trophic level (e.g. 2, 3 or 4), but rather a continuous descriptor of a 70 trophic position (e.g. 2.1, 2.7, 3.9), which can easily be calculated using a dietary matrix of 71 72 the food web. Traditionally, trophic position of a species has been evaluated by stomach content analysis, but in the past decades stable nitrogen isotopes ratios ( $\delta^{15}$ N measured as 73 the<sup>15</sup>N/<sup>14</sup>N ratio compared to a standard) has been more commonly used to assess a relative 74 trophic position of organisms. The heavier isotope <sup>15</sup>N is retained in the organism to a larger 75 extent than <sup>14</sup>N, with a relative increase of <sup>15</sup>N over <sup>14</sup>N ( $\delta^{15}$ N) of 3-5‰ per trophic level, 76 depending of species comparison and ecosystem  $^{11,12}$ . The  $\delta^{15}$ N ratios thus provide a non-77 discrete measure of the relative trophic positions along a continuum, and has been utilized in 78 ecotoxicology (either as  $\delta^{15}$ N or converted to trophic position) since the early 1990s <sup>3,5,6,13,14</sup>. 79 80

In studies of biomagnification, measurements of  $\delta^{15}N$  and contaminants are reflecting 81 82 accumulation over time. As such they are assumed to be good estimators of the average ecological (diet) and contaminant status of the respective species. Although there is increasing 83 knowledge of ecological and analytical factors that affect the variance in the contaminants 84 quantified, fewer ecotoxicological studies appreciate the unknowns and evaluate the 85 uncertainty associated to measured  $\delta^{15}N$  values, and the estimated trophic positions <sup>3</sup>. In 86 addition to a switch in diet that affect the  $\delta^{15}$ N, the isotopic ratio may vary within a species 87 depending on the productivity of the ecosystem, e.g. in phytoplankton and zooplankton the 88  $\delta^{15}$ N vary up to 5% depending on bloom stage <sup>15</sup>. This difference corresponds to a difference 89 of more than one trophic level, using the scaling factor relating relative  $\delta^{15}$ N measurements 90 with trophic levels ( $\Delta N$ )<sup>12</sup> in the range 3-5‰. Unless other information is available, a value 91

of 3.4‰ is commonly applied in ecotoxicological studies for the estimation of trophic 92 position and TMF<sup>3,16</sup>. Lastly there are analytical considerations that affect the quantified 93  $\delta^{15}$ N, such as extraction method, and removal of lipid and carbonate or not<sup>17</sup>. Using 94 95 measurements of the isotope ratios to estimate the trophic position of individuals (as opposed to estimating trophic position of a species) will make sure that some of the natural variability 96 in diets is taken into account, and this will directly affect the TMF, especially the precision. 97 98 On the other hand, it is still a model relating the individual isotope levels to trophic positions. 99 To our knowledge, no examination of the effect of variability in either enrichment factor ( $\Delta N$ ), baseline  $\delta^{15}N$  or individual sample  $\delta^{15}N$  on estimated trophic level has been performed. 100 Fortunately, ecotoxicology and risk assessment is developing in the direction of appreciating 101 and quantifying uncertainties, including an increased focus on probabilistic risk assessment 102 e.g. <sup>18,19</sup>. Thus, focus on assessing uncertainty and variability in bioaccumulation models e.g. 103 <sup>20,21</sup>, methods are needed for reducing uncertainty in TMF estimates while incorporating 104 variability in these factors. However, most TMF studies lack of appreciation of this 105 variability, i.e. most TMFs are calculated only using traditional regression methods that only 106 107 take into account (or try to minimize) error in the measured values of contaminant concentrations. Some simple methods have been performed, e.g. removing one of the 108 measured compartments from TMF calculation, as in <sup>5,22,23</sup>. Ways forward should include 109 direct quantification and treatment of the trophic level variability associated with TMF 110 111 estimates.

112

In the present study, we utilized both measurements of  $\delta^{15}N$  as well as knowledge about the structure of a food web (in the form of a binary (0/1) dietary matrix) to predict  $\delta^{15}N$  values (and hence trophic levels). The model also estimated parameters used in relating  $\delta^{15}N$  values to trophic levels (baseline/reference  $\delta^{15}N$  and enrichment factor  $\Delta N$ ), and the error variance of

117  $\delta^{15}$ N. Together the estimated parameters use the links between dietary information and 118 isotope enrichment to generate probability distributions of trophic levels, and these in turn are 119 used to generate probability distribution of TMFs.

120

#### 121 THEORY AND METHODS

122 Trophic magnification factors are assumed to reflect the magnitude of contaminant 123 accumulation in a food web, and are defined as the estimated slope of the solvent (often lipid) 124 normalized contaminant concentrations ( $C_{lipid}$ ) on trophic level (TL) (eq. 1);

$$log_{10}(C_{lipid}) = a + b \cdot TL + \varepsilon,$$
(Eq. 1)  

$$TMF = 10^{b}$$

Regressions like these are often performed by traditional least-squares regression or other 125 maximum likelihood measures attempting to minimize the squared error  $(\varepsilon)$ , i.e. the best 126 estimate of TMF are achieved through minimizing the (squared) difference between predicted 127 and observed (log) contaminant concentrations. Implicitly this means that all variability in 128 trophic levels (including measurement errors, and estimates of isotope enrichment factors etc.) 129 130 are ignored; or more correctly trophic levels are seen as independent. Though methods for 131 inclusion of errors or variability in the independent variable (so-called errors-in-variables models, e.g. Deming regression) exist, to our knowledge no such examples exist for TMF 132 estimation. Thus TMFs as measures of contaminant biomagnification does not include any 133 134 treatment of the potential variability of trophic levels among individuals or samples of the same species or population. 135

136

137 **Trophic level estimation from food webs and isotope ratio measurements.** Estimation 138 of trophic levels using  $\delta^{15}$ N is performed using equation 2:

$$TL_{consumer} = \frac{\delta^{15}N_{consumer} - \delta^{15}N_{primary\ consumer}}{\Delta N} + TL_{primary\ consumer} \quad (Eq\ 2)$$

139 where TL<sub>consumer</sub> is the trophic level of an individual with a measured  $\delta^{15}N_{consumer}$ . 140  $\delta^{15}N_{primary\ consumer}$  is the isotope ratio measured for a primary consumer assumed to occupy a 141 trophic level of TL<sub>primary\ consumer</sub>. Isotope enrichment factors ( $\Delta N$ ) of 3.4 ‰ are commonly 142 used<sup>3,16</sup>.

143

Describing the community using a food web dietary matrix yields another way to estimate 144 trophic levels. Effective trophic levels can be defined as the weighted average length of all 145 energetic pathways originating from outside a system to a specific compartment. For a 146 secondary consumer feeding on only one primary consumer this corresponds to an effective 147 trophic level of 3 (abiotic environment (TL 0)  $\rightarrow$  primary producer (TL 1)  $\rightarrow$  primary 148 consumer (TL 2)  $\rightarrow$  secondary consumer (TL 3)). With mixed diets one calculates a weighted 149 average for each compartment in the food web matrix e.g. <sup>24,25</sup>. For each species or population 150 *i* with a diet consisting of G other species according to the fraction  $F_{ii}$ , effective trophic level 151 is then calculated as: 152

$$TL_i = 1 + \sum_{j \in G} F_{ij} TL_j.$$
 (Eq 3)

$$TL = \sum (I - F)^{-1}$$
 (Eq 4)

where *I* is the identity matrix and *F* is the dietary matrix describing the food web.

By rearranging equation 2 we can use trophic levels from a dietary matrix to predictisotope ratios:

$$\delta^{15}N_i = (TL_i - TL_j)\Delta N + \delta^{15}N_j$$
 (Eq 5)

159 A Bayesian model of  $\delta^{15}$ N ratios inferred from food webs. In Bayesian statistics, the 160 goal is to arrive at distributions of parameters that reflect our degree of belief in their values. 161 The main ingredient of Bayesian analysis is Bayes rule;

$$p(\theta|y) = \frac{p(y|\theta)p(\theta)}{p(y)},$$
 (Eq 6)

where  $\theta$  represents a set of estimated parameters and y represents data or observations. Our 162 163 main goal is to get an estimate of the distribution on the left-hand-side (called a *posterior* distribution); a probability distribution of (a set of) parameters, given our data. In a simple 164 case it could be the estimate of a regression coefficient, given a sample and the distribution 165 166  $(p(\theta|y))$  could be described in terms of percentiles and a visual representation of the posterior distribution. Bayes rule gives us a way to calculate such posterior distributions since they are 167 (by definition) the product of the *likelihood* (the probability of the observations, given the 168 parameters,  $p(y|\theta)$  and a prior distribution ( $p(\theta)$ ). A likelihood is a formal measure of the 169 similarity between predictions and observations, most often directly related to sums of squares 170 and a *prior* distribution is reflecting our current knowledge about the probability of the 171 172 parameters. In the case of estimating a regression coefficient (like the TMF), we might for 173 instance have prior knowledge (from other studies or common sense) about its expected distribution. In the case of estimating the regression coefficient b in eq 1, we could form a 174 *prior distribution* which would encapsulate our current knowledge about the system, say with 175 176 a mean of 2 and a given standard deviation, if such priors were warranted based on earlier analyses. In other cases we have little information about the expected value and choose 177 178 uninformative priors, distributions that are uniform or in other ways express vague

179 information about a parameter. In most cases the likelihood,  $p(y|\theta)$ , is a combination of a 180 mathematical model that yields predictions and a model for the distribution of the errors, i.e. 181 the expected deviances between observed and predicted values. The denominator in eq 6 182 gives the probability of the observations. This is independent of the parameters of the model 183 ( $\theta$ ) and is therefore often reduced to an unknown constant yielding

$$p(\theta|y) \propto p(y|\theta)p(\theta).$$
 (Eq 7)

184 In other words, since p(y) is constant we can estimate the posterior distribution,  $p(\theta|y)$ , as 185 proportional to the prior distribution,  $p(\theta)$ , multiplied by the likelihood  $p(y|\theta)$ .

A model (Figure 1) was set up using the equations 4 and 5 to predict the population means 186 of  $\delta^{15}N$  ratios in the food web compartments, by estimating a set of parameters through 187 188 Bayesian inference. The parameters to be estimated were the non-zero entries in the dietary matrix (*F* in eq 3), the isotope enrichment factor ( $\Delta N$  in eq 5) and the population mean  $\delta^{15}$ N 189 for one of the diet matrix compartments (Daphnia, as primary consumer in eq 2). All of these 190 parameters can be combined with an error variance ( $\sigma^2$ ) estimated (common for all 191 populations) to predict  $\delta^{15}$ N in an individual (technically this error variance is a combination 192 of variance in the population and observational error). The data points of  $\delta^{15}$ N measurements 193  $(y_{ij}, i = 1, ..., n_i, j = 1, ..., J)$  are modelled as independently normally distributed within each 194 population (*j*) with means  $\mu_i$  and variance  $\sigma^2$ . The group or population means are assumed to 195 be related through the food web, according to equation 5. 196

197

198 Letting  $\theta$  denote the parameters of the dietary matrix,  $\mu_D$  the estimated population mean 199 level of  $\delta^{15}$ N for *Daphnia*,  $\sigma^2$  the variance of the  $\delta^{15}$ N distributions (common for all 200 populations) and  $\Delta N$  the isotope enrichment factor we will explore the posterior distribution

$$p(\theta, \mu_D, \Delta N, \sigma^2 | y)$$

$$\propto p(\theta, \delta^{15} N_D, \Delta N, \sigma^2) p(y|\theta, \delta^{15} N_D, \Delta N, \sigma^2)$$
(Eq 8)

201 where the likelihood is defined by:

$$p(y|\theta,\mu_{D},\Delta N,\sigma^{2}) = \prod_{j=1}^{J} \prod_{i=1}^{n_{j}} p(y_{ij}|\theta,\mu_{D},\Delta N,\sigma^{2})$$
$$= \prod_{j=1}^{J} \left(\frac{1}{2\pi\sigma^{2}}\right)^{n_{j}/2} e^{\left(\frac{\sum_{i=1}^{n_{j}} (y_{ij}-\mu_{j})^{2}}{2\sigma^{2}}\right)}.$$
(Eq 9)

In equation 9,  $y_{i,j}$  are observed isotope ratios in sample *i* belonging to population *j*, and  $\mu_j$ are the mean isotope ratios for the population *j*, given by:

$$\mu_j = f(\theta, \Delta N, \mu_D) = \left( TL_j(\theta) - TL_D(\theta) \right) \Delta N + \mu_D$$
 (Eq 10)

where  $TL_j$  is the trophic level calculated using the food web matrix as in eq 4.

205

MCMC implementation and prior probabilities. To explore the posterior values (i.e. 206 arriving at a distribution for the parameters in eq 8) we used standard Markov Chain Monte 207 Carlo (MCMC) simulations where the proposal values were generated by a normal 208 distribution around the current value <sup>26</sup>. The proposed values were accepted using the 209 Metropolis Hastings algorithm. The step size was in an initial run found so as to achieve well 210 mixed chains with an acceptance rate around 0.23 and was fixed for the main analysis <sup>26</sup>. We 211 212 simulated 10 independent chains for 100 000 iterations each and used the last 25 000 213 iterations as parameter estimates and for posterior predictive sampling. To evaluate the effect of including knowledge about the structure of the food web we also performed a Bayesian 214 215 analysis of the regression in eq 1 through Gibbs sampling, also with 10 chains for 100 000 iterations. This essentially copies the standard methods for TMF estimation<sup>3</sup>, which was also 216

applied for this specific food web<sup>23</sup>, by using a Bayesian estimation of the TMF values, while
assuming the isotope enrichment factors and all other measurements to be fixed. The analysis
was implemented in Matlab <sup>27</sup>.

220

Dirichlet distributions with concentration parameter  $\alpha = 1$  were used as priors for the diets; essentially this entails a uniform distribution over all possible combinations. Gaussian priors were used for the isotope enrichment factor ( $\Delta N$ ) and mean  $\delta^{15}N$  for *Daphnia* ( $\mu_D$ ) with means and standard deviations of (0.0035, 3×10<sup>-4</sup>) and (8, 1) respectively. For the error variance ( $\sigma^2$ ) a uniform prior with range [0...10] was applied.

226

Posterior predictive sampling and TMFs estimation. The probability distributions of 227 the estimated parameters can be used for posterior predictive sampling, essentially generating 228 distributions of  $\delta^{15}$ N values for individual samples of the different compartments in the food 229 web. For each of the  $\delta^{15}$ N data we also have contaminant data, and by resampling  $\delta^{15}$ N values 230 from the estimated distributions of  $\delta^{15}$ N we can thereby quantify the uncertainty in trophic 231 magnification factors arising from the variability in the trophic levels assigned to the analysed 232 individuals. We did this by randomly drawing *n* number of the last 25 000 iterations, using the 233 parameter values at that point in the chain to draw simulated  $\delta^{15}N$  values for the individual 234 samples (see Data sources below). Using these simulated  $\delta^{15}N$  values together with  $\Delta N$ , we 235 then performed a regression to get *n* number of estimates of TMFs for selected compounds. 236 These estimates were pooled to generate a probability distribution of TMFs given the 237 structure of the food web, the prior distributions of the parameters and observed levels of 238 contaminant and  $\delta^{15}N$ . 239

240	Data sources and food web structure. Empirical data used in the present study are
241	previously presented <sup>23,28</sup> and details on contaminant levels, sampling and analysis can be
242	found therein. In brief, representatives of the pelagic food web of Lake Mjøsa, Norway, were
243	collected mid-lake near Helgøya in September-October 2010. The food web representatives
244	included the top predator piscivorous brown trout (Salmo trutta), the zooplanktivorour fish
245	smelt (Osmerus eperlanus) and vendace (Coregonus albula). The invertebrate representatives
246	included Mysis relicta and zooplankton (Daphina galeata and Limnocalanus macrurus). The
247	samples were analysed for lipids and legacy persistent organic pollutants including
248	polychlorinated biphenyls (PCBs), polybrominated diphenyl ethers (PBDEs) and
249	dichlorodiphenyldichloroethylene $(p, p'-DDE)^{28}$ . $\delta^{15}N$ and cyclic volatile methyl siloxanes
250	(decamethylcyclopentasiloxane - D5) were analysed as described in Borgå et al. <sup>23</sup> .
251	
252	Based on previous ecological studies of Lake Mjøsa, or similar lakes, a binary dietary
253	matrix representing who eats whom (but not the proportions) for each food web representative
254	was developed. All entries in the dietary matrix were estimated; however, which entries were
255	non-zero was based on earlier studies and constitutes all the knowledge about the food web
256	included in the model. In addition to the food web compartments described above that were
257	analysed for contaminants, lipids and $\delta^{15}$ N, particulate organic matter (POM),
258	microzooplankton, small size group of vendace (< 15 cm) and smelt (< 15 cm) were included
259	in the binary dietary matrix.
260	
261	
201	DESULTS AND DISCUSSION
202	Our analysis consists of two major parts: the first part uses the assumed structure of the
205	our analysis consists of two major parts, the mist part uses the assumed structure of the

food web (i.e. who eats who), the relations in eq 4 and 5 and observations of isotope levels to

estimate the relevant parameters (diets, enrichment factor, isotope ratios for *Daphnia* and an error term) of our model. The second part uses these estimates to generate ranges of likely isotope ratios. These generated isotope ratios ( $\delta^{15}$ N), baseline isotope ratios for *Daphnia* ( $\mu_D$ ) and enrichment factor ( $\Delta N$ ) are then used to calculate trophic levels and the probability distributions of TMFs. In essence we are estimating a mean isotope ratio for each compartment and then simulating likely  $\delta^{15}$ N measurements given our model, and combining these simulated isotope ratios with observed contaminant concentrations to estimate TMFs.

273 The MCMC algorithm applied was successful in estimating the posterior distribution of diets, enrichment factor, mean isotope ratio for Daphnia and the error variance of the model. 274 275 The chains converged quickly and arrived at an acceptance rate of 0.189 during the last 25 276 000 iterations of all the 10 chains. The posterior dietary matrix (Figure 2) shows that there is quite a large range of uncertainty with regard to the feeding relations in some compartments 277 278 (especially the small smelt and vendace, and brown trout), whereas for other populations a narrower posterior was found. As  $\delta^{15}$ N values were not available for small smelt and vendace 279 (only large fish), this may explain the larger uncertainty for these compartments in the 280 posterior dietary matrix, as well as for trout that assumed to have small smelt and small 281 282 vendace as their main prey.

**Enrichment factor** –  $\Delta \delta^{15}$ **N.** The posterior for the isotope enrichment factor ( $\Delta$ N) was not very different from the prior (Figure 3), meaning that our model and observations could not adequately narrow down the distribution, thus underlining the importance of the variability in this scaling factor. For future analyses we would recommend an even wider prior range for the enrichment factor since the analysis did not narrow down the distribution substantially. The 95% credibility interval<sup>25</sup> for the enrichment factor spanned from 2.77 to 3.97 ‰ with a median of 3.29 ‰, lower than the commonly used value of 3.4 ‰. This suggests a lower

enrichment for the Mjøsa food web than previously have been assumed<sup>29</sup>. In general, the 290 291 relationship between isotope enrichment factor and TMF is such that an increase in the enrichment factor will make the estimated TMF tend away from 1. This means that assuming 292 293 a low enrichment factor will increase the risk of Type II error, i.e. increase the likelihood of classifying a magnifying compound as non-magnifying by 'pushing' the estimate towards 1. 294 295 Such issues will be even more problematic in a frequentist approach, where the main 296 questions asked is 'how probable are these contaminant observations in the food web given no 297 magnification' where non-magnifying compounds are defined as chemicals which does not exhibit a TMF significantly above 1. 298

The estimated  $\Delta N$  in our model are generally lower than the assumed value of 3.4 ‰ used in <sup>23</sup>, the probability of the enrichment factor being lower than 3.4 ‰ is 0.64 and the probability of the factor being lower than 3.0 ‰ is also substantial (0.13). This is one of the major factors that lead to our estimates of TMF being slightly lower (i.e. closer to 1) for all analysed compounds (Table 1) compared to the earlier analysis<sup>23</sup> and in the simple Bayesian regression.

The enrichment factor  $\Delta N$  is obviously associated with variability across time, space and 305 trophic level, and may be more appropriate on some specific trophic steps than others. This is 306 307 in contrast to previous studies that report one similar enrichment factor throughout the food web<sup>30</sup>, except for birds. Experimental studies on cormorants indicate that the  $\Delta N$  from bird 308 diet to muscle tissue is 2.4‰<sup>31</sup>, which is less than the recommended 3.4‰. A Bayesian 309 approach (or more generally a distributional approach) to performing analyses with  $\Delta N$  has 310 311 the possibility of including this uncertainty and quantifying it. Our model explicitly takes this uncertainty in  $\Delta N$  into account by using a distribution of the enrichment factor derived from 312 our observations and the structure of the food web. Extending this approach to include 313 distributions of  $\Delta N$  for separate groups could be valuable. 314

Predicted trophic level and TMF. One of the benefits of a Bayesian approach to 316 parameter estimation is that instead of point estimates of parameters or regression 317 318 coefficients, whole probability distributions are generated. These parameter distributions can then be used to generate more realistic predictions, since the natural variability in parameters, 319 320 such as the enrichment factor, will be included in the estimate and the generation of the 321 prediction distributions. Figure 4 show the predicted trophic levels of the populations in the food web when taking the uncertainty in diets, enrichment factor and error variance into 322 account. By using these simulated trophic levels a narrower estimate of TMFs for all 323 324 compounds are achieved, when compared to a standard Bayesian regression analysis of the observations alone (Table 1, Figure 5), despite the considerable uncertainty in some of the 325 326 parameters (e.g. the diets). Using such Bayesian approaches can also lead to a better 327 understanding of the uncertainties in bioaccumulation measures. Instead of using point estimates of TMF, as previously done in most TMF studies e.g. <sup>5,14,22</sup>, we can quantify the 328 329 uncertainty. For our model here, for instance, we can quantify the total uncertainty; given our 330 model and parameter estimates, there is a 89 % probability that the TMF for PCB-153, is greater than 2. For the cyclic siloxane D5, there is a 56% probability that the TMF is greater 331 332 than 2.

333

In summary, we have utilized Bayesian inference on the model relating relative isotope levels and trophic levels together with the structure of the food web to reduce the uncertainty in TMF estimates. With relatively few data points the method manages to estimate the diets of the species in the system, and use these diets to restrict the plausible values of trophic position of the species, and thereby also reducing the uncertainty surrounding TMF estimates. Such

reduction of uncertainty in the TMF estimate is especially of interest in cases where TMF is

close to 1, i.e. where there is a question of biomagnification, or not.

341

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- 434

437	<b>Table 1.</b> Distributions of trophic magnification factors (TMF; 2500 draws from each chain =
438	25000 simulated TMFs) for simple Bayesian regression (simple) and posterior predictive
439	simulation of the full model (full). The TMFs were determined for lipid normalized
440	concentrations. The simple model is identical to the regression model presented in the
441	empirical study <sup>22</sup> , adapted to a Bayesian framework.

MF	Model	2.5%	Median	97.5%
CB-153	Full	3.54	4.67	6.20
	Simple	3.06	4.91	7.77
CB-180	Full	4.05	5.56	7.62
	Simple	3.65	6.01	9.85
P',DDE	Full	2.99	3.8	4.87
	Simple	2.55	3.89	5.92
DE-47	Full	4.11	5.58	7.68
	Simple	3.48	5.83	9.93
DE-99	Full	1.82	2.32	3.04
	Simple	1.09	2.44	5.4
5	Full	1.66	2.03	2.45
	Simple	1.09	2.29	4.75

#### 444 FIGURE LEGENDS

445

Figure 1. Conceptual diagram of the Bayesian Food Web isotope Level Estimator. Boxes 446 447 are estimated parameters; rounded corners imply calculated values and circle represent observations. The diets of all populations (except for mikrozooplankton) are estimated. These 448 diets are used to calculate trophic levels for all compartments, using equation 4. Together with 449 450 independently estimated  $\mu_{\rm D}$  and the isotope enrichment factor ( $\Delta N$ ) these values are used to calculate isotope population means for all compartments, using equation 5. With an estimated 451 error variance these can be used to predict the observed  $\delta^{15}N$  values. For the estimation of 452 these parameters the only information used are the observed  $\delta^{15}N$  values as well as the 453 structure of the food web. 454

455

Figure 2. Estimated parameters of the food web from the Bayesian analysis. All priors used were uninformative Dirichlet distributions (i.e. uniform in *n*-dimensional space), the only previous knowledge included in the estimation was which entries in the matrix that were non-zero. Note that the distributions are highly correlated, also across compartments. X-axes are from 0 to 1, and Y is scaled to highest probability for the 51 bins used to generate the histograms.

462

Figure 3. Isotope enrichment factor (ΔN). Prior (line) and posterior (histogram)
probability distribution of the ΔN. The posterior distribution has 2.5 50 and 97.5 percentiles
of 2.77 ‰, 3.29 ‰ and 3.97 ‰.

466

467 Figure 4. Posterior predictive simulation of trophic levels for the biological
468 compartments. Lines span from 2.5 to 97.5 percentiles, bars at 25 and 75 % with diamond

469	indicating the median value. The trophic levels were simulated by selecting sets of parameters
470	from the converged chains (i.e. diets, $\mu_D$ for daphnia, $\Delta N$ and variance estimate of $\delta^{15}$ N
471	estimates). The $\delta^{15}N$ means were then calculated for all compartments and a deviation was
472	added using the variance estimation. These 'simulated' $\delta^{15}N$ values were then back calculated
473	to trophic levels using eq 2. Note that these estimates will be correlated, i.e. a higher trophic
474	level for trout is accompanied by higher trophic levels for the species in its diet. The
475	independently estimated isotope level for Daphnia used to fix the relationship in eq 5 had a
476	median value of 8.107 with a 95% confidence interval from 7.229 to 9.035.
477	
478	<b>Figure 5.</b> Distributions of trophic magnification factors (TMF; 2500 draws from each chain =
479	25000 simulated TMFs) for simple Bayesian regression (grey lines) and posterior predictive

480 simulation of the full model (black). See Table 1 for median and 95% confidence intervals.



**Figure 1.** 

	POM	Micro- zooplankton	Daphnia	Copepod	Mysis	Vendace small	Vendace large	Smelt small	Smelt large	Trout
POM										
Micro- zooplankton	1									
Daphnia										
Copepod										
Mysis										
Vendace small										
Vendace large										
Smelt small										
Smelt large										
Trout										

**Figure 2**.





**Figure 4.** 



1			
2			
3			
4			

# 6 Figure 5.

TOC / Abstract art figure: 

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